Bayesian Analysis of Simple Step-stress Model under Weibull Lifetimes ¹

Ayon Ganguly

Department of Statistics, University of Pune

Co-authors

Prof. Debasis Kundu² and Dr. Sharmishtha Mitra²

January 01, 2014

 $^{^1\}mathsf{Part}$ of this work has been supported by grants from DST and CSIR, Government of India.

Main Sections

- Censoring
- 2 Step-stress Life Tests
- 3 A Brief Literature Review
- 4 Model Description and Prior Assumptions

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- 5 Posterior Analysis
- 6 Data Analysis
 - 7 Conclusion
- 8 Future Works

Censoring

Censoring

- Quite useful technique in reliability life testing.
- Possible termination of experiment before failing all the experimental units.
- Lower cost in terms of money and time than full experiment.
- Survival experimental units can be used for further experiments.

- *n* : Number of items put on the test.
- τ : Pre-fixed time.
- $\tau^* = \tau$: Experiment termination time.



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- Number of failures is a random variable.
- Advantage : Pre-fixed experiment termination time.
- Disadvantage : Very few failures, even no failure, before time τ .

• *n* : Number of items put on the test.

- $r (\leq n)$: Pre-fixed integer.
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- Duration of experiment is a random variable.
- Advantage : Pre-fixed number of failures.
- Disadvantage : Long experimental duration.

Other Censoring Schemes

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- Progressive Censoring Schemes: Allow to remove items from the test before completion of the experiment.
- Progressive Hybrid Censoring Schemes: Mixture of hybrid and progressive censoring schemes.
- All the censoring schemes suffer form the disadvantage of either Type-I or Type-II censoring scheme.

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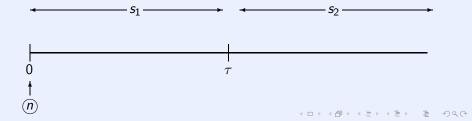
Accelerated Life Tests

- Useful experimental technique to obtain data on the lifetime distribution of highly reliable products.
- Put a sample of products on the test in some extreme environmental conditions to get early failures.
- Need to extrapolate to estimate the lifetime distribution under the normal condition.

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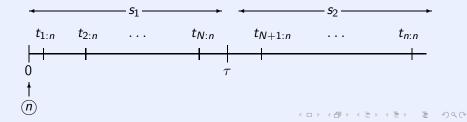
Step-stress Life Tests

- A particular type of accelerated life test.
- Allows the experimenter to change the stress levels during the life-testing experiments.
- *n* : Number of items put on the test.
- s_1, s_2 : Stress levels (Simple SSLT).
- τ : Stress changing time (Pre-fixed).



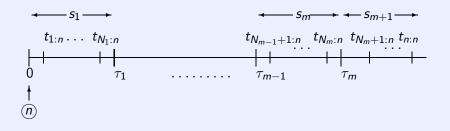
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Generalization

- n : No of items placed on the test.
- $s_1, s_2, s_3, \ldots, s_{m+1}$: Stress levels.
- $\tau_1 < \tau_2 < \ldots < \tau_m$: Stress changing times (Pre-fixed).



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Models

- Consider a simple SSLT, *i.e.*, only two stress levels, *s*₁ and *s*₂, present.
- F_i(.): CDF of lifetime of an item under the stress level s_i,
 i = 1, 2 ..., m + 1.
- F(.): CDF of life time of an item under the step-stress pattern.

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- Model needed to relate $F(\cdot)$ to $F_i(\cdot)$, i = 1, 2, ..., m+1.
- Popular models
 - Cumulative exposure model.
 - Tampered failure rate model.
 - Khamis-Higgins model.

Cumulative Exposure Model

- First proposed by Seydyakin (1966)⁴ and later studied by Nelson(1980)⁵.
- *F_i*(·) is the CDF of lifetime of an item under the stress level *s_i*, *i* = 1, 2, ..., *m* + 1.
- *F*(·) is the CDF of lifetime of an item under the step-stress pattern.

⁴Seydyakin, N. M. (1966) On one physical principle in reliability theory, *Technical Cybernatics*, 3:80-87.

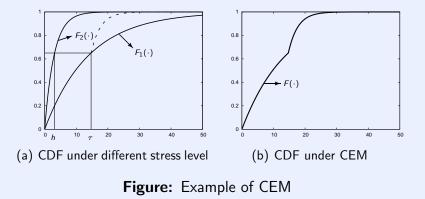
⁵Nelson (1980) Accelerated life testing: step-stress models and data analysis, *IEEE Transactions on Reliability*, 141:288-2838.

Cumulative Exposure Model

The CEM assumptions are:

- The remaining life of an item depends only on the current cumulative fraction accumulated, regardless how the fraction accumulated.
- If the stress level is fixed, the survivors will fail according to the distribution function of that stress level but starting at previous accumulated fraction failed.

Cumulative Exposure Model



Here $F_1(\cdot)$ and $F_2(\cdot)$ are CDF of Exp(14) and Exp(1) respectively.

Cumulative Exposure Model

Under the assumptions of CEM, the CDF of the lifetime is given by

$${F_{ ext{CEM}}}(t) = {F_i}(t - { au_{i-1}} + {h_{i-1}}) \;\; ext{ if } au_{i-1} \leq t < au_i, \; i = 1, \, 2, \, \ldots, \; m+1,$$

where $\tau_0 = 0$, $\tau_{m+1} = \infty$, $h_0 = 0$ and h_i , i = 1, 2, ..., m, is the solution of

$$F_{i+1}(h_i) = F_i(\tau_i - \tau_{i-1} + h_{i-1}).$$

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Tampered Failure Rate Model

- Proposed by Bhattacharyya and Soejoeti (1989)¹ for simple SSLT.
- Generalized by Madi (1993)² for multiple step SSLT.

¹Bhattacharyya, G. K. and Soejoeti, Z. (1989), A tampered failure rate model for step-stress accelerated life test, *Communication in Statistics - Theory and Methods*, 18:1627–1643.

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Tampered Failure Rate Model

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- Generalized by Madi (1993)² for multiple step SSLT.
- Effect of switching the stress level is to multiply the failure rate of the first stress level by a positive constant.

$$\lambda_{ extsf{tfrm}}(t) = \left(\prod_{j=0}^{i-1} lpha_j
ight)\lambda(t) ext{ if } au_{i-1} \leq t < au_i, \ i=1, \, 2, \, \ldots, \, m+1.$$

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Khamis-Higgins Model

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Khamis-Higgins Model

Proposed by Khamis and Higgins (1998)¹ for Weibull lifetimes.
Under KHM, the CDF is given by

$$F_{ ext{KHM}}(t) = 1 - e^{-\lambda_i (t^eta - au_{i-1}^eta) - \sum_{j=1}^{i-1} \lambda_j (au_j^eta - au_{j-1}^eta)} ext{ if } au_{i-1} \leq t < au_i.$$

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 Xu and Tang (2003)² showed that KHM is a particular case of TFRM.

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Advantages

• By increasing the stress level, reasonable number of failure can be obtained.

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• Experimental time is reduced.

2. Step-stress Life Tests

Disadvantages

- Exact relationship between the stress level and lifetime of the product is needed.
- Model must take into account the effect of stress accumulated.

• Model becomes more complicated.

A Brief Literature Review

- Balakrishnan et al. (2007)¹.
 - ► Simple step-stress life test.
 - ► Type-II censoring.
 - Exponentially distributed failure times.
 - ► Cumulative exposure model.

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$$f_{\widehat{\theta}_1}(t) = \sum_{j=1}^{J-1} \sum_{k=0}^{J} c_{jk} f_{\mathsf{G}}(t-\tau_{ik}; j, \frac{j}{\theta_1}).$$

- c_{jk} involves $(-1)^k$, $\binom{n}{j}$, $\binom{j}{k}$, and $e^{-\frac{\tau}{\theta_1}(n-j+k)}$.
- $f_{G}(\cdot)$ is the PDF of Gamma distribution.

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- ▶ MLE does not exist in explicit form.
- ► Further analysis depends on asymptotic results.

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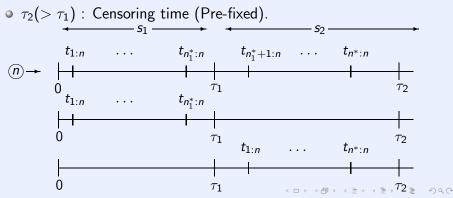
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Model Description and Prior Assumptions

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- s_1, s_2 : Stress levels.
- τ_1 : Stress changing time (Pre-fixed).
- Type-I censored data.
- $\tau_2(>\tau_1)$: Censoring time (Pre-fixed).
- Life time at stress level s_i, i = 1, 2, has a Weibull(β, λ_i) distribution, *i.e.*, its CDF is given by

$$F_i(t) = \left\{egin{array}{cc} 1-e^{-\lambda_i t^eta} & ext{if } t>0 \ 0 & ext{otherwise.} \end{array}
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• Under KHM, the CDF is given by $F_{\text{KHM}}(t) = \begin{cases} 0 & \text{if } t < 0 \\ 1 - e^{-\lambda_1 t^{\beta}} & \text{if } 0 \le t < \tau_1 \\ 1 - e^{-\lambda_2 (t^{\beta} - \tau_1^{\beta}) - \lambda_1 \tau_1^{\beta}} & \text{if } \tau_1 < t < \infty. \end{cases}$

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• Under KHM, the CDF is given by $F_{\text{KHM}}(t) = \begin{cases} 0 & \text{if } t < 0 \\ 1 - e^{-\lambda_1 t^{\beta}} & \text{if } 0 \le t < \tau_1 \\ 1 - e^{-\lambda_2 (t^{\beta} - \tau_1^{\beta}) - \lambda_1 \tau_1^{\beta}} & \text{if } \tau_1 < t < \infty. \end{cases}$

- KHM is mathematically tractable than CEM.
- It is difficult to distinguish between CEM and KHM.

Prior Assumptions I

- $\lambda_1 \sim \text{Gamma}(a_1, b_1).$
- $\lambda_2 \sim \text{Gamma}(a_2, b_2).$
- $\beta \sim \text{Gamma}(a_3, b_3).$
- λ_1 , λ_2 , and β are independently distributed.

Prior Assumptions II

- Main aim of SSLT is to get rapid failure by imposing extreme environmental condition.
- Plausible to assume that the mean life time at stress level s₂ is smaller than that at stress level s₁.

• $\lambda_1 < \lambda_2$.

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- Reparameterize $\lambda_1 = \alpha \lambda_2$ with $0 < \alpha < 1$.

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- $\lambda_2 \sim \text{Gamma}(a_2, b_2).$
- $\beta \sim \text{Gamma}(a_3, b_3)$.
- $\alpha \sim \text{Beta}(a_4, b_4)$.
- α , β , and λ_2 are independently distributed.

4. Model Description and Prior Assumptions Motivation

Motivation

• Weibull distribution is quite flexible and fits a large range of lifetime data.

Motivation

- Weibull distribution is quite flexible and fits a large range of lifetime data.
- MLEs of the model parameters do not have explicit form and all inferences rely on asymptotic distributions.

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Posterior Analysis

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Posterior Distribution under Prior Assumptions I

• For
$$\beta > 0$$
, $\lambda_1 > 0$, and $\lambda_2 > 0$
 $l_1(\beta, \lambda_1, \lambda_2 | \text{Data}) \propto \beta^{n^* + a_3 - 1} \lambda_1^{n_1^* + a_1 - 1} \lambda_2^{n_2^* + a_2 - 1} \times e^{-(b_3 - c_1)\beta - \lambda_1 A_1(\beta) - \lambda_2 A_2(\beta)},$

$$n^{*} = n_{1}^{*} + n_{2}^{*}, c_{1} = \sum_{j=1}^{n^{*}} \ln t_{j:n},$$

$$A_{1}(\beta) = b_{1} + \sum_{j=1}^{n_{1}^{*}} t_{j:n}^{\beta} + (n - n_{1}^{*})\tau_{1}^{\beta},$$

$$A_{2}(\beta) = b_{2} + \sum_{j=n_{1}^{*}+1}^{n^{*}} (t_{j:n}^{\beta} - \tau_{1}^{\beta}) + (n - n^{*})(\tau_{2}^{\beta} - \tau_{1}^{\beta}).$$

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$$n^{*} = n_{1}^{*} + n_{2}^{*}, c_{1} = \sum_{j=1}^{n^{*}} \ln t_{j:n},$$

$$A_{1}(\beta) = b_{1} + \sum_{j=1}^{n_{1}^{*}} t_{j:n}^{\beta} + (n - n_{1}^{*})\tau_{1}^{\beta},$$

$$A_{2}(\beta) = b_{2} + \sum_{j=n_{1}^{*}+1}^{n^{*}} (t_{j:n}^{\beta} - \tau_{1}^{\beta}) + (n - n^{*})(\tau_{2}^{\beta} - \tau_{1}^{\beta}).$$

• $I_1(\beta, \lambda_1, \lambda_2 | \text{Data})$ is integrable if proper priors are assumed on all the unknown parameters.

Posterior Distribution under Prior Assumptions II

• For
$$0 < \alpha < 1$$
, $\beta > 0$, and $\lambda_2 > 0$
 $l_2(\alpha, \beta, \lambda_2 | \text{Data}) \propto \alpha^{n_1^* + a_4 - 1} (1 - \alpha)^{b_4 - 1} \beta^{n^* + a_3 - 1} \lambda_2^{n^* + a_2 - 1} \times e^{-(b_3 - c_1)\beta - \lambda_2(\alpha D_1(\beta) + D_2(\beta) + b_2)},$

$$n^{*} = n_{1}^{*} + n_{2}^{*}, c_{1} = \sum_{j=1}^{n^{*}} \ln t_{j:n},$$

$$D_{1}(\beta) = \sum_{j=1}^{n_{1}^{*}} t_{j:n}^{\beta} + (n - n_{1}^{*})\tau_{1}^{\beta},$$

$$D_{2}(\beta) = \sum_{j=n_{1}^{*}+1}^{n^{*}} (t_{j:n}^{\beta} - \tau_{1}^{\beta}) + (n - n^{*})(\tau_{2}^{\beta} - \tau_{1}^{\beta}).$$

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Posterior Distribution under Prior Assumptions II

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$$0 < \alpha < 1$$
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 $l_2(\alpha, \beta, \lambda_2 | \text{Data}) \propto \alpha^{n_1^* + a_4 - 1} (1 - \alpha)^{b_4 - 1} \beta^{n^* + a_3 - 1} \lambda_2^{n^* + a_2 - 1} \times e^{-(b_3 - c_1)\beta - \lambda_2(\alpha D_1(\beta) + D_2(\beta) + b_2)},$

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5. Posterior Analysis Posterior Analysis

Bayes Estimate and Credible Interval

• Squared error loss function.

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$$g_B(\beta, \lambda_1, \lambda_2) = \int \int \int g(\beta, \lambda_1, 2) I_1(\beta, \lambda_1, \lambda_2) d\lambda_2 d\lambda_1 d\alpha.$$

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- Squared error loss function.
- $g_B(\beta, \lambda_1, \lambda_2) = \int \int \int g(\beta, \lambda_1, 2) l_1(\beta, \lambda_1, \lambda_2) d\lambda_2 d\lambda_1 d\alpha.$
- Bayes estimate of g(β, λ₁, λ₂) cannot be obtained explicitly in general.

- Squared error loss function.
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- Bayes estimate of g(β, λ₁, λ₂) cannot be obtained explicitly in general.
- An algorithm based on importance sampling is proposed to compute ĝ_B(β, λ₁, λ₂) and to construct CRI for g(β, λ₁, λ₂) in both the cases.

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$$l_1(\beta, \lambda_1, \lambda_2 \mid \text{Data}) = l_3(\lambda_1, \mid \beta, \text{Data}) \times l_4(\lambda_2, \mid \beta, \text{Data}) \times l_5(\beta \mid \text{Data}),$$

where

$$\begin{split} I_{3}(\lambda_{1}, | \beta, \text{Data}) &= \frac{\{A_{1}(\beta)\}^{n_{1}^{*}+a_{1}}}{\Gamma(n_{1}^{*}+a_{1})} \lambda_{1}^{n_{1}^{*}+a_{1}-1} e^{-\lambda_{1}A_{1}(\beta)} \quad \text{if} \quad \lambda_{1} > 0, \\ I_{4}(\lambda_{2}, | \beta, \text{Data}) &= \frac{\{A_{2}(\beta)\}^{n_{2}^{*}+a_{2}}}{\Gamma(n_{2}^{*}+a_{2})} \lambda_{2}^{n_{2}^{*}+a_{2}-1} e^{-\lambda_{2}A_{2}(\beta)} \quad \text{if} \quad \lambda_{2} > 0, \\ I_{5}(\beta | \text{Data}) &= c_{2} \frac{\beta^{n^{*}+a_{3}-1}e^{-(b_{3}-c_{1})\beta}}{\{A_{1}(\beta)\}^{n_{1}^{*}+a_{1}}\{A_{2}(\beta)\}^{n_{2}^{*}+a_{2}}} \quad \text{if} \quad \beta > 0. \end{split}$$

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Data Analysis

6. Data Analysis

Illustrative Example

• An artificial data is generated from KHM with n = 40, $\beta = 2$, $\lambda_1 = 1/1.2 \simeq 0.833$, $\lambda_2 = 1/4.5 \simeq 2.222$, and $\tau_1 = 0.6$.

6. Data Analysis

Illustrative Example

An artificial data is generated from KHM with n = 40, β = 2, λ₁ = 1/1.2 ⊆ 0.833, λ₂ = 1/4.5 ⊆ 2.222, and τ₁ = 0.6.
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- $a_1 = b_1 = a_2 = b_2 = a_3 = b_3 = 0.0001$ and $a_4 = b_4 = 1$.
- Prior I : $\hat{\beta} = 2.35$, $\hat{\lambda}_1 = 0.93$, $\hat{\lambda}_2 = 2.61$.
- Prior II : $\hat{\beta} = 2.49$, $\hat{\lambda}_1 = 1.01$, $\hat{\lambda}_2 = 2.50$.

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- Prior II : $\widehat{\beta} = 2.49$, $\widehat{\lambda}_1 = 1.01$, $\widehat{\lambda}_2 = 2.50$.
- Prior I : 95% symmetric CRI for β is (1.12, 4.04).
- Prior II : 95% symmetric CRI for β is (0.45, 2.44).

Conclusions

 Extensive simulation has been done to judge the performance of the proposed procedures.

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 MSEs of BE of all unknown parameters are smaller in case of Prior II than those in case of Prior I.

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- MSEs of BE of all unknown parameters are smaller in case of Prior II than those in case of Prior I.
- Other loss functions and other censoring schemes can be handled in a very similar fashion.

Future Works

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8. Future Works

Future Works

• Step-stress model in the presence of competing risks under Bayesian framework.

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Future Works

• Step-stress model in the presence of competing risks under Bayesian framework.

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• Optimality of SSLT under Bayesian framework.

Future Works

- Step-stress model in the presence of competing risks under Bayesian framework.
- Optimality of SSLT under Bayesian framework.
- Prior elicitation is becoming a popular topic among Bayesian. It will be a challenging task to find a subjective prior for step-stress life testing models.

Thank You

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